

Alice HH vs Bob HT using an Edgeworth Expansion

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1 Definitions

Symbol	Definition
$n \in \mathbb{N}$	Number of coinflips in the sequence
$C_1 C_2, \dots, C_n$	Sequence of coinflips $C_i \in \{H, T\}$
$\Delta_i := 1_{C_i C_{i+1}=HH} - 1_{C_i C_{i+1}=HT}$	The difference (Alice score - Bob score) due to the coinflips C_i, C_{i+1}
$X_n = \sum_{i=1}^{n-1} \Delta_i$	Alices total Score minus Bob's total Score. The total difference in score after the n flips
$1_{HH\bullet}$ and $1_{\bullet HH}$	Used as a shorthand notation for the indicators of the events like $1_{HH\bullet} := 1_{\{C_i C_{i+1} C_{i+2} = HH\bullet\}}$ when the value of i is implicit. The bullet “ \bullet ” stands for a “wildcard” that could be either H or T , but makes calculations a bit easier to visualize when used in multiple places to keep i constant. For example: $1_{HH\bullet} \cdot 1_{\bullet HH} = 1_{HHH}$ or $1_{HH\bullet} \cdot 1_{\bullet HT} = 1_{HHT}$ or $1_{HT\bullet} \cdot 1_{\bullet HT} = 0$ (since the middle coinflip cannot be both H and T simultaneously).

2 Moment Calculations

2.1 1st Moment

$$\mathbf{E}[X_n] = (n-1)\mathbf{E}[\Delta_i] = 0$$

2.2 2nd Moment

Note that Δ_i and Δ_j are independent unless $i = j$ or $i = j \pm 1$ (i.e. the “range” of the interactions is only 1 coinflip.) From this it follows that we can expand to get:

$$\mathbf{E}[X_n^2] = (n-1)\mathbf{E}[\Delta_i^2] + 2(n-2)\mathbf{E}[\Delta_i \Delta_{i+1}]$$

and we compute:

$$\mathbf{E}[\Delta_i^2] = \mathbf{E}[(1_{HH} - 1_{HT})^2] = \mathbf{E}[1_{HH} + 1_{HT} + 2 \cdot 0] = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$$

and:

$$\begin{aligned} \mathbf{E}[\Delta_i \Delta_{i+1}] &= \mathbf{E}[(1_{HH\bullet} - 1_{HT\bullet})(1_{\bullet HH} - 1_{\bullet HT})] \\ &= \mathbf{E}[1_{HHH} - 1_{HHT} - 0 + 0] \\ &= \frac{1}{16} - \frac{1}{16} = 0 \end{aligned}$$

so combining gives:

$$\mathbf{E}[X_n^2] = \frac{1}{2}(n-1)$$

2.3 3rd Moment

Again, keeping track of only interactions at range 1 away, we have the expansion:

$$\begin{aligned}\mathbf{E}[X_n^3] &= \mathbf{E}\left[\sum_{i,j,k=1}^n \Delta_i \Delta_j \Delta_k\right] \\ &= (n-1)\mathbf{E}[\Delta_i^3] + 3(n-3) \cdot 3! \cdot \mathbf{E}[\Delta_i \Delta_{i+1} \Delta_{i+2}] \\ &\quad + (n-2) \cdot \binom{3}{1} \mathbf{E}[\Delta_i^2 \Delta_{i+1}] \\ &\quad + (n-2) \cdot \binom{3}{1} \mathbf{E}[\Delta_i \Delta_{i+1}^2]\end{aligned}$$

Now we notice that any expectation of the form $\mathbf{E}[(\dots \text{product of stuff} \dots) \cdot \Delta_{i+2}]$ (i.e. it ends with a Δ on its own and not squared) is exactly 0 because the the last Δ is equally likely to be +1 or -1 (either both 50% or both 0% depending on whether the product of stuff at the beginning end is a H or ends in a T.) So the only surviving term here is the term $\mathbf{E}[\Delta_i \Delta_{i+1}^2]$ which gives:

$$\begin{aligned}\mathbf{E}[\Delta_i \Delta_{i+1}^2] &= \mathbf{E}\left[(1_{HH\bullet} - 1_{HT\bullet})(1_{\bullet HH} - 1_{\bullet HT})^2\right] \\ &= \mathbf{E}[(1_{HH\bullet} - 1_{HT\bullet})(1_{\bullet HH} + 1_{\bullet HT} + 2 \cdot 0)] \\ &= \mathbf{E}[1_{HHH} + 1_{HHT} - 0 - 0] \\ &= \frac{1}{8} + \frac{1}{8} = \frac{1}{4}\end{aligned}$$

So we have:

$$\mathbf{E}[X_n^3] = \frac{3}{4}(n-2)$$

3 Edgeworth Expansion

We use an expansion for the density function of the random variable of the form (see https://en.wikipedia.org/wiki/Edgeworth_series)

$$\rho_X(x) \approx \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \left(1 + \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x-\mu}{\sigma}\right) + \dots\right)$$

where $\mu = \mathbf{E}[X]$, $\sigma = \sqrt{\mathbf{Var}[X]}$ and $\kappa_3 = \mathbf{E}[(X - \mu)^3]$, and $He_3(x) = x^3 - 3x$ is the 3rd Hermite polynomial. In our case $\mu = 0$ so this simplifies a bit to:

$$\rho_X(x) \approx \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left(1 + \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x}{\sigma}\right) + \dots\right)$$

3.1 A lemma about integrating He_3

Lemma 1. *Have that:*

$$\int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} He_3\left(\frac{x}{\sigma}\right) dx = -\frac{1}{\sqrt{2\pi}}$$

Proof. Have:

$$\begin{aligned}
\int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} He_3\left(\frac{x}{\sigma}\right) dx &= \mathbf{E}_{Z \sim \mathcal{N}(0, \sigma^2)} \left[\mathbf{1}_{Z>0} He_3\left(\frac{Z}{\sigma}\right) \right] \\
&= \mathbf{E}_{Z \sim \mathcal{N}(0,1)} [\mathbf{1}_{Z>0} He_3(Z)] \\
&= \mathbf{E}_{Z \sim \mathcal{N}(0,1)} [\mathbf{1}_{Z>0} Z^3] - 3\mathbf{E}_{Z \sim \mathcal{N}(0,1)} [\mathbf{1}_{Z>0} Z] \\
&= \mathbf{E}_{Z \sim \mathcal{N}(0,1)} [\varphi(Z)^3] - 3\mathbf{E}_{Z \sim \mathcal{N}(0,1)} [\varphi(Z)] \\
&= (3-1)!! \frac{1}{\sqrt{2\pi}} - 3 \frac{1}{\sqrt{2\pi}} \\
&= -\frac{1}{\sqrt{2\pi}}
\end{aligned}$$

where we have used the relu function $\varphi(x) = x\mathbf{1}_{\{x>0\}}$ and the result for Gaussians that (which can be proved by a nice integration by parts induction)

$$\mathbf{E} [\varphi(Z)^k] = \begin{cases} \frac{(k-1)!!}{\sqrt{2\pi}} & k \text{ is even} \\ \frac{(k-1)!!}{\sqrt{2\pi}} & k \text{ is odd} \end{cases}$$

□

3.2 Alice - Bob using the Edgeworth Approximation.

Lemma 2. *If we use the Edgeworth approximation:*

$$\rho_X(x) \approx \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left(1 + \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x}{\sigma}\right) + \dots \right)$$

then we get:

$$\mathbf{P}(X > 0) - \mathbf{P}(X < 0) = -\frac{2\kappa_3}{3!\sigma^3} \frac{1}{\sqrt{2\pi}}$$

Proof. By making the change of variable $x \rightarrow -x$ and using the fact that $He_3(x)$ is an odd polynomial:

$$\begin{aligned}
\mathbf{P}(X < 0) &= \int_{-\infty}^0 \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left(1 + \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x}{\sigma}\right) + \dots \right) dx \\
&= \int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left(1 + \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{-x}{\sigma}\right) + \dots \right) dx \\
&= \int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left(1 - \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x}{\sigma}\right) + \dots \right) dx
\end{aligned}$$

Hence, when we subtract the probabilities we get:

$$\begin{aligned}
\mathbf{P}(X > 0) - \mathbf{P}(X < 0) &= \int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} \left((1-1) + (1-(-1)) \frac{\kappa_3}{3!\sigma^3} He_3\left(\frac{x}{\sigma}\right) + \dots \right) dx \\
&= 2 \frac{\kappa_3}{3!\sigma^3} \cdot \int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} He_3\left(\frac{x}{\sigma}\right) dx + \dots \\
&= -\frac{2\kappa_3}{3!\sigma^3} \frac{1}{\sqrt{2\pi}}
\end{aligned}$$

by the Lemma.

□

Proposition 3. Suppose that the edgeworth approximation holds for the random variable X_n which is Alice's Score minus Bob's score. Then:

$$\mathbf{P}(\text{Alice wins}) - \mathbf{P}(\text{Bob wins}) = -\frac{1}{2\sqrt{\pi n}} + \dots$$

Remark 4. Note that we know already from other methods that the result of the proposition is true. However, it is **NOT YET PROVEN** that the Edgeworth expansion actually holds (need to deal with the fact that the sequence Δ_i are not purely independent but are instead dependent)

Proof. By the lemmas we have:

$$\begin{aligned} \mathbf{P}(\text{Alice wins}) - \mathbf{P}(\text{Bob wins}) &= \mathbf{P}(X_n > 0) - \mathbf{P}(X_n < 0) \\ &= -\frac{2\kappa_3}{3!\sigma^3} \frac{1}{\sqrt{2\pi}} + \dots \\ &= -\frac{2\left(\frac{3}{4}(n-2)\right)}{3!\left(\frac{1}{2}(n-1)\right)^{3/2}} \frac{1}{\sqrt{2\pi}} + \dots \\ &= -\frac{2\left(\frac{3}{4}\right)}{3!\left(\frac{1}{2}\right)} \frac{1}{\sqrt{\frac{1}{2}}} \frac{(n-2)}{\sqrt{2\pi}(n-1)^{3/2}} + \dots \\ &= -\frac{1}{2\sqrt{\pi n}} + \dots \end{aligned}$$

where we have used $\frac{n-2}{(n-1)^{3/2}} = \frac{1}{\sqrt{n}} + \dots$ as $n \rightarrow \infty$. □