

# On the Discrepancy of Random Matrices with Many Columns

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Cole Franks and Michael Saks

September 27, 2018



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- Formally, let  $\|\cdot\|_*$  be a norm, and let

$$\text{disc}_*(M) = \min_{v \in \{+1, -1\}^n} \|Mv\|_*$$

( $M$  is an  $m \times n$  matrix).

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## Examples and Applications

$$\text{disc}_\infty \begin{bmatrix} 1 & 0 & 1 \\ 1 & 1 & 1 \end{bmatrix} = 1$$

- Extractors: the best extractor for two independent  $n$ -bit sources with min-entropy  $k$  has error rate  $\text{disc}_\infty(M)$  where  $M$  is a
  1.  $\binom{2^n}{2^k}^2 \times 2^{2n}$  matrix
  2. with one row for each rectangle  $A \times B \subset \{0, 1\}^n \times \{0, 1\}^n$  with  $|A| = |B| = 2^k$ ,
  3. each row is a  $2^n \times 2^n$  matrix with  $(x, y)$  entry equal to  $\frac{1}{2^{2k}} 1_A(x) 1_B(y)$ .

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# Upper bounds

## Definition

$\text{herdisc}(M)$ : maximum discrepancy of any subset of columns of  $M$ .

Beck-Fiala Theorem:  $M_{ij} \in [-1, 1]$  and  $\leq t$  nonzero entries per column,

$$\text{herdisc}(M) \leq 2t - 1.$$

Beck-Fiala Conjecture: If  $M$  as above,

$$\text{herdisc}(M) = O(\sqrt{t})$$

Komlos Conjecture:  $M$  with unit vector columns,

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# Discrepancy of random matrices

Let  $M$  be a random  $t$ -sparse matrix

$$m \underbrace{\begin{bmatrix} 0 & 1 & 1 & 0 & 1 \\ 1 & 0 & 0 & 1 & 1 \\ 1 & 1 & 1 & 1 & 0 \end{bmatrix}}_n$$

Theorem (Ezra, Lovett 2015)

Few columns: If  $n = O(m)$ , then with probability  $1 - \exp(-\Omega(t))$ .

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Many columns: If  $n = \Omega\left(\binom{m}{t} \log \binom{m}{t}\right)$  then with pr.  $1 - \binom{m}{t}^{-\Omega(1)}$ ,

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## General setup

- $\mathcal{L} \subset \mathbb{R}^m$  is a nondegenerate lattice,
- $X$  is a finitely supported r.v. on  $\mathcal{L}$  such that  $\text{span}_{\mathbb{Z}} X = \mathcal{L}$ .
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For  $n \gg m$  the problem becomes a **closest vector** problem on  $\mathcal{L}$ .

## Definition

$\rho_*(\mathcal{L})$  is the covering radius of  $\mathcal{L}$  in the norm  $\|\cdot\|_*$ .

## Fact

$\text{disc}_*(M) \leq 2\rho_*(\mathcal{L})$  with high probability as  $n \rightarrow \infty$ .

Naïvely,  $n$  has to be huge.

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- $\mathcal{L}$  is  $\{x \in \mathbb{Z}^m : \sum x_i \equiv 0 \pmod t\}$
- $\rho_\infty(\mathcal{L}) = 1$

By fact,  $\text{disc}_\infty(M) \leq 2$  eventually.

EL15 showed this happens for  $n \geq \Omega\left(\binom{m}{t} \log \binom{m}{t}\right)$ . exponential dependence on  $t!$

This work:  $n = \Omega(m^3 \log^2 m)$

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Random  $t$ -sparse matrices:

## Theorem (FS18)

Let  $M$  be a random  $t$ -sparse matrix. If  $n = \Omega(m^3 \log^2 m)$ , then

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with probability at least  $1 - O\left(\sqrt{\frac{m \log n}{n}}\right)$ .

Actually usually  $\text{disc}_\infty(M) = 1$ .

**Related work:** Hoberg and Rothvoss '18 obtained  $\Omega(m^2 \log m)$  for  $M$  with i.i.d  $\{0, 1\}$  entries.

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## More generally

$\mathcal{L}, M, X$  as before, and define

1.  $L = \max_{v \in \text{supp } X} \|v\|_2$

e.g.  $\sqrt{t}$  for  $t$ -sparse

2. distortion  $R_* = \max_{\|v\|_2=1} \|v\|_*$ .

e.g.  $\sqrt{m}$  for  $* = \infty$

3. spanningness:  $s(X)$  “how far  $X$  is from proper sublattice.”

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### Theorem (FS18)

Suppose  $\mathbb{E}XX^\dagger = I_m$ . Then  $\text{disc}_*(M) \leq 2\rho_*(\mathcal{L})$  with probability  $1 - O\left(L\sqrt{\frac{\log n}{n}}\right)$  for

$$n \geq N = \text{poly}(m, s(X)^{-1}, R_*, \rho_*(\mathcal{L}), \log \det \mathcal{L}).$$

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Suppose  $\mathbb{E}XX^\dagger = I_m$ . Then  $\text{disc}_*(M) \leq 2\rho_*(\mathcal{L})$  with probability  $1 - O\left(L\sqrt{\frac{\log n}{n}}\right)$  for

$$n \geq N = \text{poly}(m, s(X)^{-1}, R_*, \rho_*(\mathcal{L}), \log \det \mathcal{L}).$$

## More generally

$\mathcal{L}, M, X$  as before, and define

1.  $L = \max_{v \in \text{supp } X} \|v\|_2$

e.g.  $\sqrt{t}$  for  $t$ -sparse

2. *distortion*  $R_* = \max_{\|v\|_2=1} \|v\|_*$ .

e.g.  $\sqrt{m}$  for  $* = \infty$

3. *spanningness*:  $s(X)$  “how far  $X$  is from proper sublattice.”

will be  $\leq 1/m$  for  $t$ -sparse

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To apply the theorem to non-isotropic  $X$ ,  
consider the transformed r.v.  $\Sigma^{-1/2}X$ , where  $\Sigma = \mathbb{E}XX^\dagger$ .

# Proof outline

Need to show: for *most fixed*  $M$ , the r.v.  $My$ ,  $y \in_R \{\pm 1\}^n$ , gets within  $2\rho_*(\mathcal{L})$  of the origin with positive probability.

Use local central limit theorem:

1. Intuitively the  $My$  (sampled at same time) approaches lattice Gaussian:

$$\Pr[My = \lambda] \propto e^{-\frac{1}{2}\lambda^\top \Sigma^{-1} \lambda}$$

for  $\lambda \in M\mathbf{1} + 2\mathcal{L}$

2. For *most*  $M$ ,  $My$  also behaves like this!
3. Then done:  $\lambda \in M\mathbf{1} + 2\mathcal{L}$  contains, near origin, elements of  $\ast$ -norm  $2\rho_*(\mathcal{L})$ .

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# Local central limit theorem

We propose an LCLT that takes a matrix parameter  $M$ , and show it holds for **most**  $M$ .

- Proof of LCLT  $\approx$  proof of LCLT in [Kuperberg, Lovett, Peled, '12].
- Differences:
  - theirs was for **FIXED** *very wide* matrices.
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## Obstruction to LCLTs:

If  $X$  lies on a proper sublattice  $\mathcal{L}' \subsetneq \mathcal{L}$ , in trouble.

Need an *approximate* version of the assumption that this doesn't happen.

# Spanningness

## Definition

Dual lattice:  $\mathcal{L}^* := \{\boldsymbol{\theta} : \forall \boldsymbol{\lambda} \in \mathcal{L}, \langle \boldsymbol{\lambda}, \boldsymbol{\theta} \rangle \in \mathbb{Z}\}$ .

## Definition

$f_X(\boldsymbol{\theta}) := \sqrt{\mathbb{E}[|\langle X, \boldsymbol{\theta} \rangle \bmod 1|^2]}$ , where  $\bmod 1 \rightarrow [-1/2, 1/2)$

$$f_X(\boldsymbol{\theta}) = 0 \implies \boldsymbol{\theta} \in \mathcal{L}^*.$$

$$f_X(\boldsymbol{\theta}) \approx 0 \implies \langle X, \boldsymbol{\theta} \rangle \approx \mathbb{Z}.$$

Thus, obstruction is  $\boldsymbol{\theta}$  far from  $\mathcal{L}^*$  with  $f_X(\boldsymbol{\theta})$  small.

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Say  $\theta$  is pseudodual if

$$f_X(\theta) \leq \frac{1}{2}d(\theta, \mathcal{L}^*).$$

(Why pseudodual? Near  $\mathcal{L}^*$ ,  $f_X(\theta) \approx d(\theta, \mathcal{L}^*$ .)

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$$G_M(\lambda) = \frac{2^{m/2} \det(\mathcal{L})}{\pi^{m/2} \sqrt{\det(MM^\dagger)}} e^{-2\lambda^\dagger(MM^\dagger)^{-1}\lambda}$$

on  $\mathbb{R}^m$  (Gaussian with covariance  $\frac{1}{2}MM^\dagger$ ).

### Theorem (FS18)

With probability  $1 - O\left(L\sqrt{\frac{\log n}{n}}\right)$  over the choice of  $M$ ,

- $\frac{1}{2}nI_m \preceq MM^\dagger \preceq 2nI_m$

- $$\left| \Pr_{y_i \in \{\pm 1/2\}} [My = \lambda] - G_M(\lambda) \right| = G_M(0) \cdot O\left(\frac{m^2 L^2}{n}\right)$$

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# Proof of local limit theorem

---

## Definition (Fourier transform!)

If  $Y$  is a random variable on  $\mathbb{R}^m$ ,  $\widehat{Y} : \mathbb{R}^m \rightarrow \mathbb{C}$  is

$$\widehat{Y}(\boldsymbol{\theta}) = \mathbb{E}[e^{2\pi i \langle Y, \boldsymbol{\theta} \rangle}].$$

## Fact (Fourier inversion:)

*if  $Y$  takes values on  $\mathcal{L}$ , then*

$$\Pr(Y = \boldsymbol{\lambda}) = \det(\mathcal{L}) \int_D \widehat{Y}(\boldsymbol{\theta}) e^{-2\pi i \langle \boldsymbol{\lambda}, \boldsymbol{\theta} \rangle} d\boldsymbol{\theta}$$

*Here  $D$  is any fundamental domain of the dual lattice  $\mathcal{L}^*$ .*

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## Take Fourier transform

For fixed  $M$ , Fourier transform of  $M\mathbf{y}$  for  $\mathbf{y} \in_R \{\pm 1/2\}$ ?

Say  $i^{\text{th}}$  column is  $\mathbf{x}_i$ .

$$\begin{aligned}\widehat{M\mathbf{y}}(\boldsymbol{\theta}) &= \mathbb{E}_{\mathbf{y}} \left[ e^{2\pi i \langle \sum_{j=1}^n y_j \mathbf{x}_j, \boldsymbol{\theta} \rangle} \right] \\ &= \prod_{j=1}^n \mathbb{E}_{y_j} [e^{2\pi i y_j \langle \mathbf{x}_j, \boldsymbol{\theta} \rangle}] \\ &= \prod_{j=1}^n \cos(\pi \langle \mathbf{x}_j, \boldsymbol{\theta} \rangle).\end{aligned}$$

## Use Fourier inversion

Let  $\varepsilon > 0$ , to be picked with hindsight (think  $n^{-1/4}$ )

$$\left| \frac{1}{\det \mathcal{L}} \Pr(My = \boldsymbol{\lambda}) - G_M(\boldsymbol{\lambda}) \right| = \left| \int_D e^{-2\pi i \langle \boldsymbol{\lambda}, \boldsymbol{\theta} \rangle} (\widehat{My}(\boldsymbol{\theta}) - \widehat{G}_M(\boldsymbol{\theta})) d\boldsymbol{\theta} \right|$$
$$\leq \int_{B(\varepsilon)} |\widehat{My}(\boldsymbol{\theta}) - \widehat{G}_M(\boldsymbol{\theta})| d\boldsymbol{\theta} \quad (I_1)$$

$$+ \int_{\mathbb{R}^m \setminus B(\varepsilon)} |\widehat{G}_M(\boldsymbol{\theta})| d\boldsymbol{\theta} \quad (I_2)$$

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If  $D \subset B(\varepsilon)$ .  $D$  is the Voronoi cell in  $\mathcal{L}^*$ .

rest of the proof is to show these are small!

- First two easy from the eigenvalue property.
- $\mathbb{E}_M[I_3] \leq e^{-\varepsilon^2 n}$  if  $\varepsilon \leq s(X)$ .

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## Applying the main theorem

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# Random $t$ -sparse matrices

From now on we just want to bound the spanningness. We'll do it for  $t$ -sparse vectors - the framework is that of [KLP12].

## Lemma

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Recall what  $s(X) \geq \frac{1}{m}$  means. We need to show that if  $\theta$  is **pseudodual**, i.e.,  $f_X(\theta) \leq \|\theta\|/2$  but *not dual*, then  $f_X(\theta) \geq \alpha/m$ .

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**Proof outline:** (recall  $f_X(\theta) := \sqrt{\mathbb{E}[|\langle X, \theta \rangle \bmod 1|^2]}$ )

- if all  $|\langle x, \theta \rangle \bmod 1| \leq 1/4$  for all  $x \in \text{supp } X$ , then  $f_X(\theta) \leq 1/4$ , so  $\theta$  not pseudodual unless dual.
- $X$  is  $\frac{1}{2m}$ -**spreading**: for all  $\theta$ ,

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# Open problems

- Can the colorings guaranteed by our theorems be produced efficiently? The probability a random coloring is good decreases with  $n$  as  $\sqrt{n}^{-m}$ , which is not good enough.
- As a function of  $m$ , how many columns are required such that  $\text{disc}(M) \leq 2$  for  $t$ -sparse vectors with high probability?

Thank you!