# Joint distribution

### Math 477

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# 1 Discrete RVs

#### 1.1 Joint distribution

Let X, Y be random variables. To have all the information about X, Y we need to know P(X = i, Y = j) for all possible i, j in the range of X, Y. However, it may NOT be that X, Y are independent. In other words, knowing P(X = i) and P(Y = j) is not enough to determine P(X = i, Y = j). In this case, we need to list all possible pari of P(X = i, Y = j) in a 2 dimensional table. This is the so-called joint distribution of X and Y.

**Example 1.1.** Suppose 3 balls are randomly selected from an urn containing 3 red, 4 white and 5 blue. Let X and Y denote, respectively, the number of red and white balls in the sample. Find P(X = i, Y = j) for all possible (i, j).

Ans: It's best to present these informations in a table. We'll just list the answers here.

$$\begin{split} p(0,0) &= \frac{\binom{5}{3}}{\binom{12}{3}}; p(0,1) = \frac{\binom{4}{1}\binom{5}{2}}{\binom{12}{3}}; p(0,2) = \frac{\binom{4}{2}\binom{5}{1}}{\binom{12}{3}}; p(0,3) = \frac{\binom{4}{3}}{\binom{12}{3}}\\ p(1,0) &= \frac{\binom{3}{1}\binom{5}{2}}{\binom{12}{3}}; p(1,1) = \frac{\binom{3}{1}\binom{4}{1}\binom{5}{1}}{\binom{12}{3}}; p(1,2) = \frac{\binom{3}{1}\binom{4}{2}}{\binom{12}{3}}\\ p(2,0) &= \frac{\binom{3}{2}\binom{5}{1}}{\binom{12}{3}}; p(2,1) = \frac{\binom{3}{2}\binom{4}{1}}{\binom{12}{3}}\\ p(3,0) &= \frac{\binom{3}{3}}{\binom{12}{3}}. \end{split}$$

For a joint distribution, one should check that

$$\sum_{i,j} P(X=i,Y=j) = 1.$$

Note that the above is a double sum over i, j.

### 1.2 Marginal distribution

Intuitively, one can recover the information about X from knowing both about X, Y. This can be done, via the following:

$$P(X = i) = sum_j P(X = i, Y = j)$$

since

$$\{X = i\} = \bigcup_{j} \{X = i, Y = j\}.$$

Similarly, we have

$$P(Y = j) = sum_i P(X = i, Y = j).$$

The distribution P(X = i) is referred to as the marginal distribution of X. The distribution P(Y = j) is referred to as the marginal distribution of Y.

# 2 Continuous random variables

# 2.1 Joint density

Taking the motivation from the definition of a continuous RV, we say two RVs X and Y are *jointly continuous* if there exists a function f(x, y) such that

$$f(x,y) \geq \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dx dy = 1.$$

0

We refer to f(x, y) as the joint density of X, Y. Then we define

$$P(x_1 \le X \le x_2, y_1 \le Y \le y_2) = \int_{x_1}^{x_2} \int_{y_1}^{y_2} f(x, y) dx dy.$$

From which it follows that

$$P(X = x, Y = y) = 0,$$

and thus

$$P(x_1 \le X \le x_2, y_1 \le Y \le y_2) = P(x_1 < X \le x_2, y_1 \le Y \le y_2)$$
  
=  $P(x_1 < X < x_2, y_1 \le Y \le y_2)$   
=  $P(x_1 \le X \le x_2, y_1 < Y \le y_2)$   
=  $P(x_1 \le X \le x_2, y_1 < Y \le y_2)$ 

as well as

$$P(X \le x, Y \le y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u, v) du dv.$$

The function  $F(x, y) = P(X \le x, Y \le y)$  is referred to as the joint cumulative distribution function of X and Y. If X, Y are continuous, then F is jointly differentiable in x, y and

$$f(x,y) = \frac{\partial^2}{\partial x \partial y} F(x,y).$$

Example 2.1. (Uniform on a unit square)

Let X, Y have a joint density

$$f(x,y) = 1, 0 \le x \le 1, 0 \le y \le 1$$
  
= 0 othersiwe.

Then for  $0 \le x \le 1, 0 \le y \le 1$ ,

$$F(x,y) = \int_0^x \int_0^y dx dy = xy.$$

Note that F(x,y) is a constant in x for  $x \ge 1$  since for  $x \ge 1$  and  $0 \le y \le 1$ 

$$F(x,y) = \int_0^1 \int_0^y dx dy = y,$$

agreeing with the fact that

$$f(x,y) = \frac{\partial^2}{\partial x \partial y} F(x,y) = 0 \text{ for } x \ge 1.$$

Similarly F(x,y) is a constant in y for  $x \ge 1$ . And finally for  $0 \le x, y \le 1$ 

$$\frac{\partial^2}{\partial x \partial y} F(x, y) = \frac{\partial^2}{\partial x \partial y} xy = 1 = f(x, y).$$

### 2.2 Marginal distribution, marginal density

Again, we have information about X, Y. How can we find information about X given this information? First, note that

$$P(X \le x) = P(X \le x, Y < \infty) = \lim_{y \to \infty} F(x, y),$$

since the set  $\{Y < \infty\}$  has probability 1. That is

$$P(X \le x) = \int_{-\infty}^{x} \int_{-\infty}^{\infty} f(u, v) du dv$$
$$= \int_{-\infty}^{x} \int_{-\infty}^{\infty} f(u, v) dv du.$$

For a fix u, note that

$$\int_{-\infty}^{\infty} f(u,v) dv$$

is a function of u. Therefore, differentiating with respect to x on both sides gives

$$f_X(x) = \int_{-\infty}^{\infty} f(x, v) dv$$

We call  $f_X(x)$  the marginal density of X and  $F_X(x) = P(X \le x)$  the marginal cdf of X. Similarly we have

$$f_Y(y) = \int_{-\infty}^{\infty} f(u, y) du,$$

and the cumulative distribution function of Y is

$$P(Y \le y) = \int_{-\infty}^{y} f_Y(u) du = \int_{-\infty}^{\infty} f(u, v) du dv.$$

### 3 Independence

We say, generally that 2 RVs X, Y are independent if for any subsets A, B of the real line:

$$P(X \in A, Y \in B) = P(X \in A)P(Y \in B).$$

In the case that X, Y are either jointly discrete or jointly continuous, we can say more.

#### 3.1 Discrete RV

Let X, Y be discrete RVs with joint distribution P(X = i, Y = j). Then X, Y are independent if and only if

$$P(X = i, Y = j) = P(X = i)P(Y = j).$$

Proof.

$$\begin{split} P(X \in A, Y \in B) &= \sum_{x \in A} \sum_{y \in B} P(X = x, Y = y) \\ &= \sum_{x \in A} \sum_{y \in B} P(X = x) P(Y = y) = \sum_{x \in A} P(X = x) \sum_{y \in B} P(Y = y) \\ &= P(X \in A) P(Y \in B). \end{split}$$

**Example 3.1.** Suppose that n + m independent trials having a common probability success p are performed. If X is the number of successes in the first n trials, Y the number of successes in m trials, then X and Y are independent. However, let Z be the number of total successes. Then X and Z are NOT independent.

### 3.2 Continuous RV

Let X, Y be discrete RVs with joint density  $f_{XY}(x, y)$ . Then X, Y are independent if and only if

$$f_{XY}(x,y) = f_X(x)f_Y(y).$$

Proof.

$$P(X \in A, Y \in B) = \int_{A} \int_{B} f_{XY}(x, y) dx dy$$
  
= 
$$\int_{A} \int_{B} f_{X}(x) f_{Y}(y) dx dy = \int_{A} f_{X}(x) \int_{B} f_{Y}(y)$$
  
= 
$$P(X \in A) P(Y \in B).$$

**Example 3.2.** Two persons decide to meet at a certain location. If each of them independently arrives at a time uniformly distributed between 12 - 1 pm, find the probability that the first to arrive has to wait longer than 10 minutes.

Ans: Let X, Y denote the time the first and the second person arrives. Then X, Y are independent Uniform(0,60). We want to compute

$$P(X + 10 < Y) + P(Y + 10 < X) = 2P(X + 10 < Y),$$

by symmetry. We have

$$2P(X+10 < Y) = 2 \iint_{x+10 < y} f(x,y) dx dy$$
  
=  $2 \int_{10}^{60} \int_{0}^{y-10} (1/60)^2 dx dy$   
=  $\frac{2}{60^2} \int_{10}^{60} (y-10) dy$   
=  $25/36.$ 

**Example 3.3.** If the join density function of X and Y is

$$f(x,y) = 6e^{-2x}e^{-3y}, 0 < x < \infty, 0 < y < \infty,$$
  
= 0 otherwise

are they independent? What if

$$f(x,y) = 24xy, 0 < x, y < 1, 0 < x + y < 1$$
  
= 0 otherwise

Ans: The RVs are independent in the first case and not in the second. The reason is if we denote

$$\mathbf{1}(x,y) = 1$$
 if  $0 < x, y < 1, 0 < x + y < 1$   
= 0 otherwise

then we see that for the second case

$$f(x,y) = 24x\mathbf{1}(x,y),$$

and clearly the function  $\mathbf{1}(x, y)$  does not factor.