

# Functions of continuous RV

Math 477

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## 1 Monotone function of a continuous RV

**Theorem 1.1.** *Let  $X$  be a continuous RV with pdf  $f_X(x)$ . Suppose  $g(x)$  is a strictly monotonic and differentiable function of  $x$ . Then  $Y = g(X)$  is a continuous RV with density*

$$\begin{aligned} f_Y(y) &= f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| \text{ if } y = g(x) \text{ for some } x \\ &= 0 \text{ otherwise.} \end{aligned}$$

*Proof.*

Suppose  $g$  is increasing.

$$P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y)) = F_X(g^{-1}(y)).$$

Differentiating yields

$$f_Y(y) = f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y).$$

The case when  $g$  is decreasing can be proved similarly.

### 1.1 Examples

**Example 1.2.** *Let  $X$  be a continuous non-negative RV with density function  $f$  and let  $Y = X^n$ . Then*

$$f_Y(y) = \frac{1}{n} y^{1/n-1} f_X(y^{1/n}).$$

**Example 1.3.** *The log normal distribution* Let  $X$  be a  $\text{Normal}(\mu, \sigma^2)$  distribution. Then  $Y = e^X$  is said to have a lognormal distribution with parameters  $\mu, \sigma^2$ .  $Y$  has density

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2 y}} e^{-\frac{(\log(y)-\mu)^2}{2\sigma^2}}.$$

## 2 Non-monotone function of continuous RV

**Example 2.1.** *Let  $X$  have Uniform  $[-1, 1]$  distribution. What is the pdf of  $Y = X^2$ ?*

Ans: For  $0 \leq y \leq 1$

$$P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = \sqrt{y}.$$

Therefore,  $f_Y(y) = \frac{1}{2\sqrt{y}}$ .

Remark: The function  $g(x) = x^2$  is NOT monotone on  $[-1, 1]$ . However, noting that for  $x < 0$ , the inverse of  $x^2$  would be  $-\sqrt{x}$  and

$$|d/dx(-\sqrt{x})| = |d/dx\sqrt{x}| = \frac{1}{2\sqrt{x}},$$

we can try to apply the Theorem (1.1) to see what happens. Specifically, we would have

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| = \frac{1}{2} \frac{d}{dy} \sqrt{y} = \frac{1}{4\sqrt{y}},$$

and this is incorrect.

Note that we can try to correct the situation by saying since there are “two” inverse functions of  $x^2$  on the two intervals  $[-1, 0]$  and  $[0, 1]$  we should add over these to get

$$f_Y(y) = \frac{1}{4\sqrt{y}} + \frac{1}{4\sqrt{y}} = \frac{1}{2\sqrt{y}},$$

which now agrees with our previous result. But note how this is a not straightforward argument.

The above correction may rely on the fact that  $f_X$  is symmetric over 0. Let’s try a non-symmetric example

**Example 2.2.** *Let  $X$  have Uniform  $[-2, 1]$  distribution. What is the pdf of  $Y = X^2$ ?*

Ans: We need to distinguish between  $1 \leq y \leq 4$  and  $0 \leq y \leq 1$ . For  $0 \leq y \leq 1$

$$P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = \frac{2\sqrt{y}}{3}.$$

For  $1 \leq y \leq 4$

$$P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq 1) = \frac{1 + \sqrt{y}}{3}.$$

Thus we see that the pdf is

$$\begin{aligned} f_Y(y) &= \frac{1}{3\sqrt{y}}, 0 \leq y < 1 \\ &= \frac{1}{6\sqrt{y}}, 1 \leq y \leq 4. \end{aligned}$$

Again, one can apply the theorem (1.1) with care. On the interval  $[-2,-1]$ , there is “one” inverse function for  $g(x) = x^2$  and on the interval  $[-1,1]$  there are “two” inverse functions for  $g(x) = x^2$ . Thus we need to add for those values of  $y$  that have the inverse in the interval  $[-1,1]$ , namely  $0 \leq y \leq 1$  giving

$$f_Y(y) = \frac{1}{3} \frac{1}{2\sqrt{y}} + \frac{1}{3} \frac{1}{2\sqrt{y}} = \frac{1}{3\sqrt{y}}.$$

And for those values of  $y$  that have inverse in the interval  $[-2, -1]$ , namely  $2 \leq y \leq 4$ , we do not have to add, giving

$$f_Y(y) = \frac{1}{3} \frac{1}{2\sqrt{y}} = \frac{1}{6\sqrt{y}}.$$

Let’s look at a last example where the inverses are different over different intervals

**Example 2.3.** Let  $X$  have Uniform  $[-1, 1]$  distribution. Define

$$\begin{aligned} g(x) &= x^2, 0 \leq x \leq 1 \\ &= x^4, -1 \leq x < 0. \end{aligned}$$

What is the pdf of  $Y = g(X)$ ?

Ans: For  $0 \leq y \leq 1$

$$P(Y \leq y) = P(-y^{1/4} \leq X \leq \sqrt{y}) = \frac{\sqrt{y} + y^{1/4}}{2}.$$

Thus,

$$\begin{aligned}f_Y(y) &= \frac{1}{4\sqrt{y}} + \frac{1}{8y^{3/4}}, 0 \leq y \leq 1 \\ &= 0 \text{ otherwise .}\end{aligned}$$

Again, note that  $g(x)$  have “two” inverses, depending on  $x \in [-1, 0]$  or  $x \in [0, 1]$ . Therefore, we need to add over these. Thus applying Theorem (1.1) gives

$$f_Y(y) = \frac{1}{2}|d/dy - y^{1/4}| + \frac{1}{2}|d/dy\sqrt{y}| = \frac{1}{4\sqrt{y}} + \frac{1}{8y^{3/4}},$$

agreeing with our previous result.

### 3 Translating and scaling of continuous RV

A particular function of interest for us, which is always monotonic is  $f(x) = \frac{x-a}{b}$ . You should verify the following results:

1. If  $X$  has a Uniform  $[a,b]$  distribution then  $Y = \frac{X-\mu}{\sigma}$  has Uniform  $[\frac{a-\mu}{\sigma}, \frac{b-\mu}{\sigma}]$  distribution.
2. If  $X$  has Normal( $\mu, \sigma^2$ ) distribution then  $Y = \frac{X-a}{b}$  has Normal( $\mu - a, (\frac{\sigma}{b})^2$ ) distribution.
3. If  $X$  has Exponential( $\lambda$ ) distribution. Then  $Y = \frac{X}{\sigma}$  have Exponential( $\frac{\lambda}{\sigma}$ ) distribution.
4. If  $X$  has Exponential( $\lambda$ ) distribution. Then  $Y = \frac{X-\mu}{\sigma}$  does NOT have Exponential distribution, simply because the support of  $Y$  no longer is  $[0, \infty)$ . However, you can still say that  $X$  has some “exponential type” distribution with support on  $[-\mu, \infty)$  and you can work out its density.