## Solutions for Homework 7, Math 477, Fall 2018

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## From the Problems in Chapter 6:

14 Let X denote the position of the ambulance at the time of the accident; X is uniformly distributed on [0, L]. Let Y denote the position of the accident; it is also uniformly distributed on [0, L]. We are asked for the distribution of the distance between the accident and the ambulance; i.e., |X - Y|.

Define U = Y - X and V = X. We seek the distribution of |U|, but let's first find the distribution of U. We consider the change of variables

$$\mathbf{U}(x,y) = (y-x,x)$$

which has the inverse

$$\mathbf{X}(u,v) = (v,u+v)$$

The joint density of (U, V) is 0 unless  $0 \le x(u, v), y(u, v) \le L$  which is the same as

$$0 \le v \le L$$
 and  $-v \le u \le L - v$ ,

and this region is the parallelepiped  $\Omega$  bounded by v = 0, v = L, u = -v and u = L - v.

In the interior of this region, the joint density of (U.V) is given by

$$\frac{1}{L^2} \left| \frac{\partial(x,y)}{\partial(u,v)} \right| = \frac{1}{L^2} \ ,$$

The probability density function of U is then the u marginal of the joint probability density function

$$g(u,v) = \begin{cases} L^{-2} & (u,v) \in \Omega \\ 0 & (u,v) \notin \Omega \end{cases}$$

The probability density of U is  $g_U(u)$ , the marginal g(u, v), which is  $\int_{\mathbb{R}} g(u, v) dv$ . This is simply  $L^{-2}$  times the length of the segment that is the intersection of the vertical line though (0, u) and  $\Omega$ . That is,

$$g_U(u) = \begin{cases} L^{-2}(L+u) & u \in [-L,0] \\ L^{-2}(L-u) & u \in [0,L] \\ 0 & u \notin [-L,L] \end{cases}$$

Finally,  $|U| \leq a$  if and only if  $-a \leq U \leq a$ , and so, for  $a \in [0, L]$ 

$$P(|U| \le a) = \int_{-a}^{a} g_U(u) du = 2 \int_{0}^{a} (1-u) du = L^{-2}(2La - a^2).$$

Evidently, for  $a \ge 2$ ,  $P(U| \le a) = 1$ .

Alternatively, one could use convolutions: X - Y = X + (-Y), and since X and -Y are independent, and -Y is uniformly distributed on [-L, 0], the convolution formula gives another way to arrive at the formula for  $g_U$ .

20 Since

$$f(x,y) = \begin{cases} xe^{-(x+y)} & x,y \ge 0\\ 0 & \text{otherwise} \end{cases}$$

where

$$f_X(x) = \begin{cases} xe^{-x} & x \ge 0\\ 0 & \text{otherwise} \end{cases} \quad \text{and} \quad f_Y(y) = \begin{cases} e^{-y} & y \ge 0\\ 0 & \text{otherwise} \end{cases}$$

X and Y are independent.

In the second case,

$$f(x,y) = \begin{cases} 2 & 0 \le x \le y \ , 0 \le y \ , 0 \le 1 \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

X and Y are not independent since the joint density function os not a product function: Indeed, since f(x, y) = 0 unless 0 < x < y, once you know, say  $Y \le 1/2$ , you know  $X \le 1/2$ ., while without being given this information about Y, one knows only that  $X \le 1$ .

**32** Let  $X_j$  denote the sales in month j. We are given that  $X_j$  has mean  $\mu = 100$  and variance  $\sigma^2 = 25$ , and that these normal variables are independent. For each j,  $P(X_j > 100) = \frac{1}{2}$ . Hence the probability that in exactly 3 of the next 6 months sales exceed 100 is the same as the probability that in tossing a fair coin 6 times, we get exactly three heads, and this probability is

$$\binom{6}{3}2^{-6} = \frac{5}{16} \; .$$

For the second part, let  $Z = X_1 + X_2 + X_3 + X_4$ . Then X is normal with mean 100 and variance 100. Hence (Z - 400)/10 is standard normal, and

$$P(Z > 420) = P\left(\frac{Z - 400}{10} > 2\right) = \Phi(2)$$
.

## From the theoretical exercises in Chapter 6:

8 Recall that the hazard rate function of a continuous random variable X is given by

$$\lambda(t) = -\frac{\mathrm{d}}{\mathrm{d}t} \ln(P(X > t)) \; .$$

Let X and Y be independent and let  $W := \min\{X, Y\}$ . Then

$$P(W > t) = P(\{X < t\} \cap \{Y > t\}) = P(X > t)P(Y > t) .$$

Thus

$$\overline{F}_W(t) = \overline{F}_X(t)\overline{F}_Y(t) \ .$$

Next, taking the logarithmic derivative,

$$\lambda_W(t) = -\frac{\mathrm{d}}{\mathrm{d}t}\ln(P(W > t)) = -\frac{\mathrm{d}}{\mathrm{d}t}\left(\ln P(X > t) + \ln P(Y > y)\right) = \lambda_X(t) + \lambda_Y(t) \;.$$